

Statistical approaches for sparse radio interferometric imaging

Error estimation in radio interferometry imaging
or
"Bayesian compressive sensing"

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Outline

- 1 Bayesian interpretations of interferometric imaging techniques
- 2 Proximal MCMC sampling
- 3 Sparse regularisation with Bayesian credible regions

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Radio interferometric inverse problem

- Consider the ill-posed inverse problem of radio interferometric imaging:

$$y = \Phi x + n,$$

where y are the measured visibilities, Φ is the linear measurement operator, x is the underlying image and n is instrumental noise.

- Measurement operator, e.g. $\Phi = \text{GFA}$, may incorporate:
 - primary beam A of the telescope;
 - Fourier transform F ;
 - convolutional de-gridding G to interpolate to continuous uv -coordinates;
 - direction-dependent effects (DDEs)...

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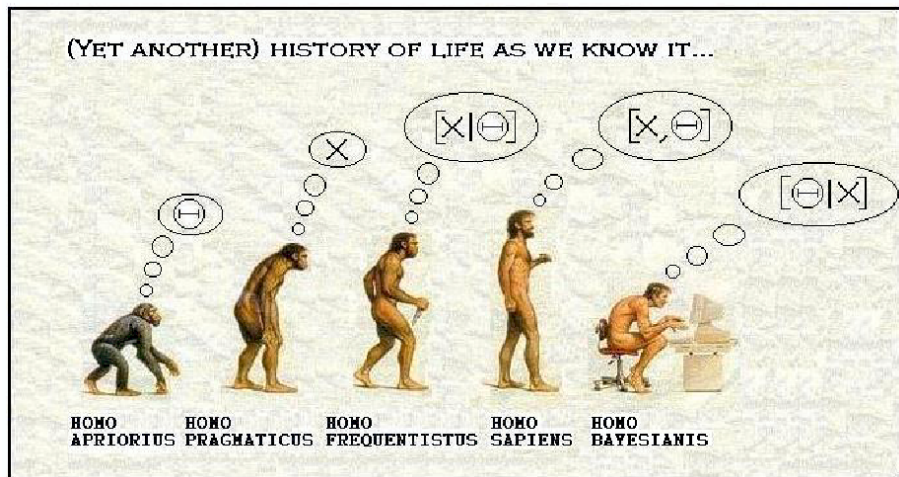
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 - direction-dependent effects (DDEs)...

Interferometric imaging: recover an image from noisy and incomplete Fourier measurements.

Bayesian evolution



Bayesian inference

- Given data \mathbf{y} (visibilities) and model M (interferometric telescope with Gaussian noise), we want a **full probabilistic description** of our knowledge of the underlying **sky image \mathbf{x}** .

- Bayes to the rescue:

$$P(\mathbf{x} | \mathbf{y}, M) = \frac{P(\mathbf{y} | \mathbf{x}, M) P(\mathbf{x} | M)}{P(\mathbf{y} | M)}$$

Bayes Theorem



- Bayes theorem in words:

$$\text{posterior} = \frac{\text{likelihood} \times \text{prior}}{\text{evidence}}$$

- How do we perform Bayesian inference in practice?
 - \Rightarrow maximum a-posteriori (MAP) estimates and sampling approaches (MCMC)
 - (and many others)

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Bayes in practice

MAP and MCMC sampling

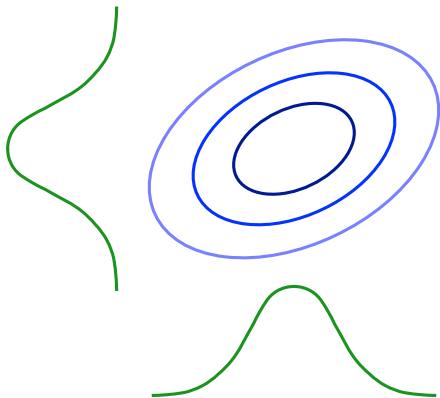


Figure: Probability distribution to explore in 2D

Bayes in practice

MAP and MCMC sampling

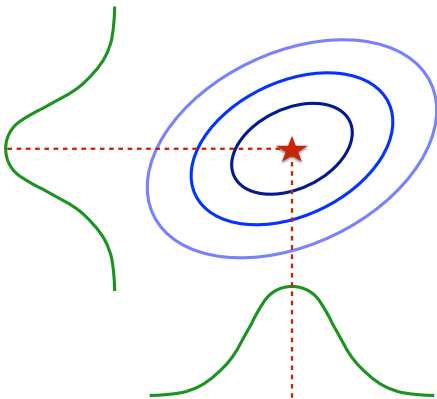


Figure: Maximum a-posteriori (MAP) estimate

Bayes in practice

MAP and MCMC sampling

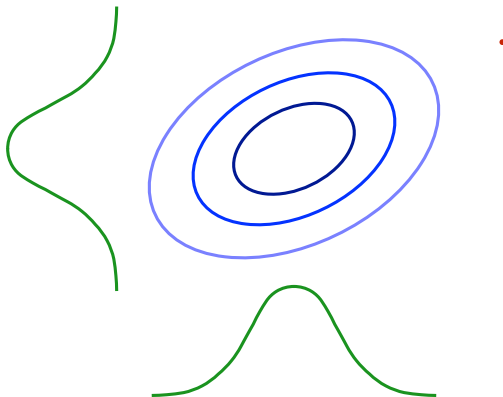


Figure: Markov Chain Monte Carlo (MCMC) sampling

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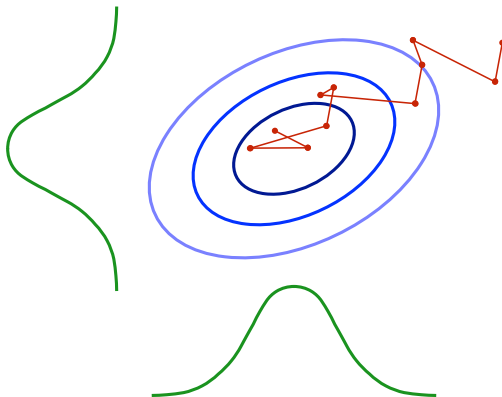


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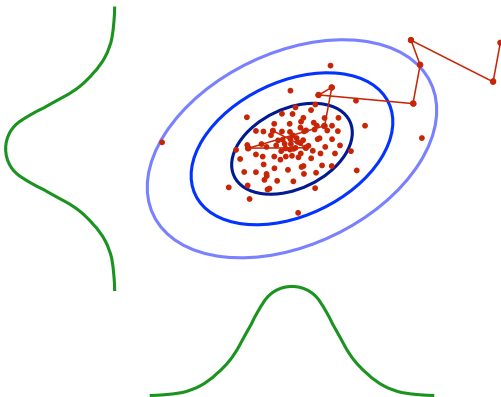


Figure: Markov Chain Monte Carlo (MCMC) sampling

MAP estimation and regularisation

Hint: they're the same thing!

- Many interferometric imaging approaches are based on [regularisation](#) (*i.e.* minimising an objective function comprised of a data-fidelity penalty and a regularisation penalty).
- Consider the MAP estimation problem. . .

MAP estimation and regularisation

Hint: they're the same thing!

- Start with Bayes Theorem (ignore normalising evidence):

$$P(\mathbf{x} | \mathbf{y}) \propto P(\mathbf{y} | \mathbf{x})P(\mathbf{x}), \quad \text{i.e. posterior} \propto \text{likelihood} \times \text{prior}$$

- MAP estimator:

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Define likelihood (assuming Gaussian noise) and prior:

$$P(\mathbf{y} | \mathbf{x}) \propto \exp\left(-\|\mathbf{y} - \Phi\mathbf{x}\|_2^2 / (2\sigma^2)\right)$$

Likelihood

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$$P(\mathbf{x}) \propto \exp\left(-R(\mathbf{x})\right)$$

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$$\log P(\mathbf{x} | \mathbf{y}) = -\|\mathbf{y} - \Phi\mathbf{x}\|_2^2 / (2\sigma^2) - R(\mathbf{x}) + \text{const.}$$

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CLEAN and MEM as MAP estimators

• CLEAN

Consider the sparse prior: $P(\mathbf{x}) \propto \exp(-\beta \|\mathbf{x}\|_0)$.

Corresponding MAP estimator is:

$$\mathbf{x}_{\text{clean}} = \arg \min_{\mathbf{x}} \left[\|\mathbf{y} - \Phi \mathbf{x}\|_2^2 + \lambda \|\mathbf{x}\|_0 \right]$$

(Laplace prior $P(\mathbf{x}) \propto \exp(-\beta \|\mathbf{x}\|_1)$ is good proxy; Wiaux *et al.* 2009)

• MEM

Consider the entropic prior: $P(\mathbf{x}) \propto \exp(-\beta \mathbf{x}^\dagger \log \mathbf{x})$.

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(In practice some differences: CLEAN does not solve MAP problem exactly;
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Compressive sensing

Synthesis framework

- Consider sparsifying representation (e.g. wavelet basis):

$$\mathbf{x} = \sum_i \Psi_i \alpha_i = \begin{pmatrix} | \\ \Psi_0 \\ | \end{pmatrix} \alpha_0 + \begin{pmatrix} | \\ \Psi_1 \\ | \end{pmatrix} \alpha_1 + \dots \Rightarrow \boxed{\mathbf{x} = \Psi \alpha}$$

- Recover (wavelet) coefficients α of image \mathbf{x} .
- Consider the Laplacian prior on coefficients: $P(\alpha) \propto \exp(-\beta \|\alpha\|_1)$.
- Sparse synthesis regularisation problem:

$$\mathbf{x}_{\text{synthesis}} = \Psi \times \arg \min_{\alpha} \left[\|\mathbf{y} - \Phi \Psi \alpha\|_2^2 + \lambda \|\alpha\|_1 \right]$$

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More sophisticated approaches

- Overcomplete dictionary composed of a concatenation of orthonormal bases:

$$\Psi = [\Psi_1, \Psi_2, \dots, \Psi_q]$$

- Constrained vs unconstrained problems.
- Re-weighted versions and the SARA algorithm (Carrillo, McEwen, Wiaux 2012).
- Sparse analysis regularisation problem:

$$\mathbf{x}_{\text{analysis}} = \arg \min_{\mathbf{x}} \left[\|\mathbf{y} - \Phi \mathbf{x}\|_2^2 + \lambda \|\Psi^\dagger \mathbf{x}\|_1 \right]$$

Analysis framework

- Analysis problem viewed from a synthesis perspective:

$$\mathbf{x}_{\text{analysis}} = \Psi \times \arg \min_{\gamma \in \text{column space}(\Psi^\dagger)} \left[\|\mathbf{y} - \Phi \Psi^\dagger \gamma\|_2^2 + \lambda \|\gamma\|_1 \right]$$

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Analysis as synthesis

Sampling the full posterior distribution

Markov Chain Monte Carlo (MCMC)

- Alternative is to **sample full posterior** distribution $P(\mathbf{x} | \mathbf{y})$.
- Provides **uncertainty (error) information**.
- MCMC methods for high-dimensional problems (like interferometric imaging):
 - Gibbs sampling (sample from conditional distributions)
 - Hamiltonian MC (HMC) sampling (exploit gradients)
 - Metropolis adjusted Langevin algorithm (MALA) sampling (exploit gradients)
- Gibbs sampling applied to radio interferometric imaging (Sutter, Wandelt, McEwen, *et al.* 2014), using methods developed for CMB by Wandelt *et al.* (2005).
 - Assume isotropic Gaussian process prior characterised by power spectrum C_ℓ .
 - Sample from conditional distributions:

$$\mathbf{x}^{i+1} \leftarrow P(\mathbf{x} | C_\ell^i, \mathbf{y}) \quad \text{and} \quad C_\ell^{i+1} \leftarrow P(C_\ell | \mathbf{x}^{i+1}).$$

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Require MCMC approach to support sparse priors, which shown to be highly effective.

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MCMC sampling with gradients

Langevin dynamics

- Work done by **Xiaohao Cai** in collaboration with **Marcelo Pereyra**.
- Consider posteriors of the following form (and more compact notation):

$$P(\boldsymbol{x} | \boldsymbol{y}) = \underbrace{\pi(\boldsymbol{x})}_{\text{Posterior}} \propto \exp\left[-\underbrace{g(\boldsymbol{x})}_{\text{Convex}}\right]$$

- If $g(\boldsymbol{x})$ differentiable can adopt MALA (Langevin dynamics) or HMC (Hamiltonian dynamics) MCMC methods.
- **Langevin dynamics** model **molecular dynamics** (includes friction and occasional high velocity collisions that perturb the system).
- Based on **Langevin diffusion process** $\mathcal{L}(t)$, with π as stationary distribution:

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where \mathcal{W} is Brownian motion.

- Need gradients so **cannot support sparse priors**.

MCMC sampling with gradients

Langevin dynamics

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- Consider posteriors of the following form (and more compact notation):

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Proximity operators

A brief aside

- Define **proximity operator**:

$$\text{prox}_g^\lambda(\mathbf{x}) = \arg \min_{\mathbf{u}} \left[g(\mathbf{u}) + \|\mathbf{u} - \mathbf{x}\|^2 / 2\lambda \right]$$

- Generalisation of **projection operator**:

$$\mathcal{P}_{\mathcal{C}}(\mathbf{x}) = \arg \min_{\mathbf{u}} \left[\iota_{\mathcal{C}}(\mathbf{u}) + \|\mathbf{u} - \mathbf{x}\|^2 / 2 \right],$$

where $\iota_{\mathcal{C}}(\mathbf{u}) = \infty$ if $\mathbf{u} \notin \mathcal{C}$ and zero otherwise.

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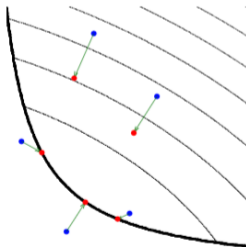


Figure: Illustration of proximity operator [Credit: Parikh & Boyd (2013)]

Proximal MCMC

Moreau approximation

- Follow Pereyra (2016a) and consider Moreau approximation of π :

$$\pi_\lambda(\mathbf{x}) = \sup_{\mathbf{u} \in \mathbb{R}^N} \pi(\mathbf{u}) \exp\left(-\frac{\|\mathbf{u} - \mathbf{x}\|^2}{2\lambda}\right)$$

- Important properties of $\pi_\lambda(\mathbf{x})$:

- As $\lambda \rightarrow 0$, $\pi_\lambda(\mathbf{x}) \rightarrow \pi(\mathbf{x})$
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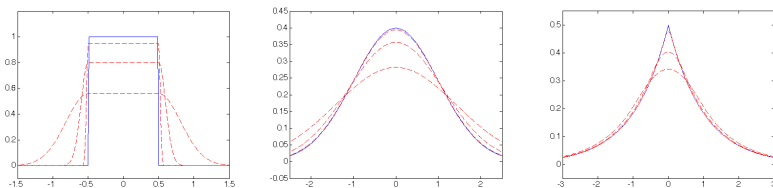


Figure: Illustration of Moreau approximations [Credit: Pereyra (2016a)]

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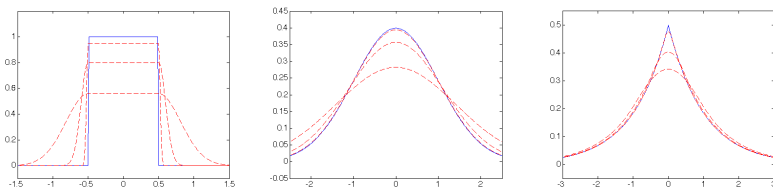


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Proximal MCMC

P-MALA

- Discretise Langevin differential equation using forward Euler approximation:

$$\mathbf{l}^{(m+1)} = \mathbf{l}^{(m)} + \frac{\delta}{2} \nabla \log \pi(\mathbf{l}^{(m)}) + \sqrt{\delta} \mathbf{w}^{(m)},$$

where δ controls the discrete-time increment and $\mathbf{w}^{(m)}$ is (unit) Gaussian distributed.

- Apply Moreau approximation and compute gradient by prox:

$$\mathbf{l}^{(m+1)} = \text{prox}_g^{\delta/2}(\mathbf{l}^{(m)}) + \sqrt{\delta} \mathbf{w}^{(m)}.$$

- (After adding a Metropolis-Hastings accept-reject step, get P-MALA with $\mathbf{l}^{(m)} \rightarrow \pi$.)
- Must compute $\text{prox}_g^{\delta/2}(\cdot)$: develop analytic expression that can be computed rapidly.

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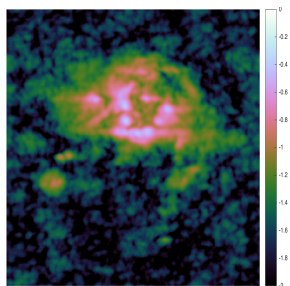
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Proximal MCMC

Preliminary results on simulations

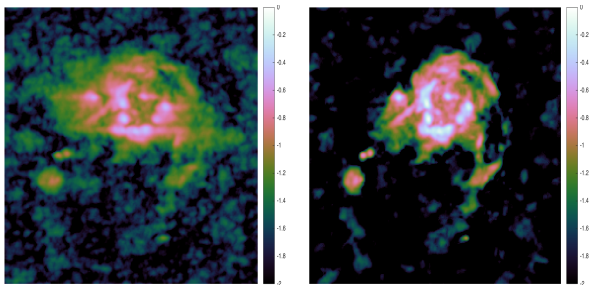


(a) Dirty image

Figure: HII region of M31

Proximal MCMC

Preliminary results on simulations



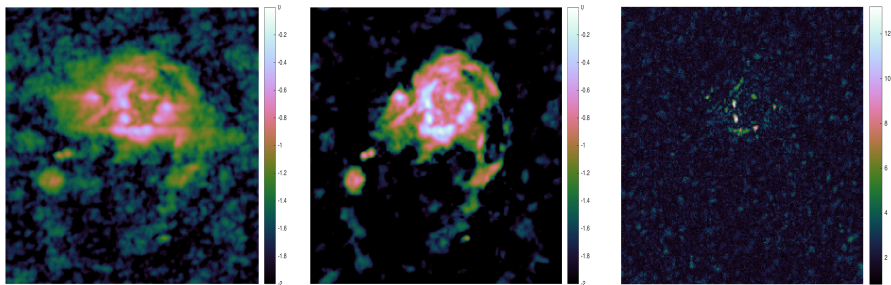
(a) Dirty image

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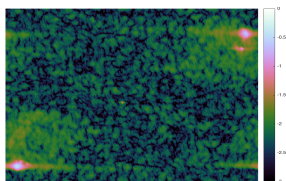
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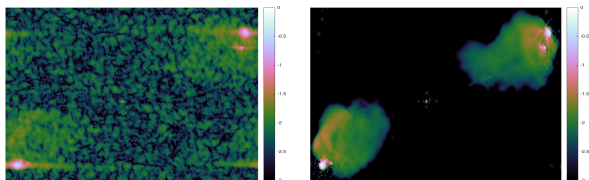


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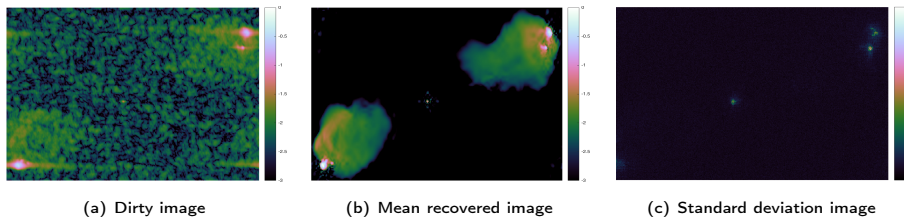
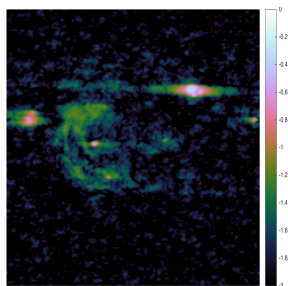


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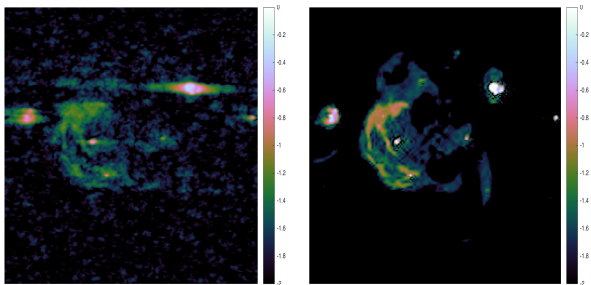


(a) Dirty image

Figure: Supernova remnant W28

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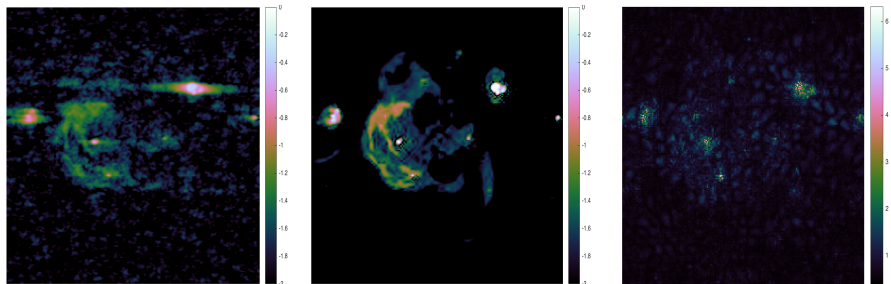
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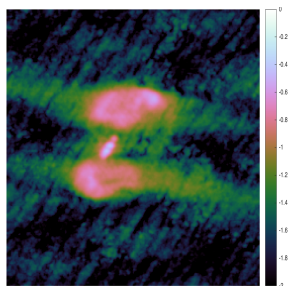
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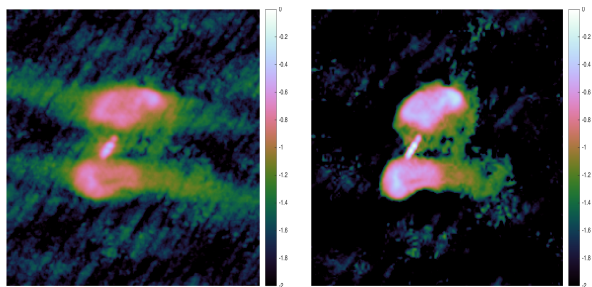


(a) Dirty image

Figure: 3C288

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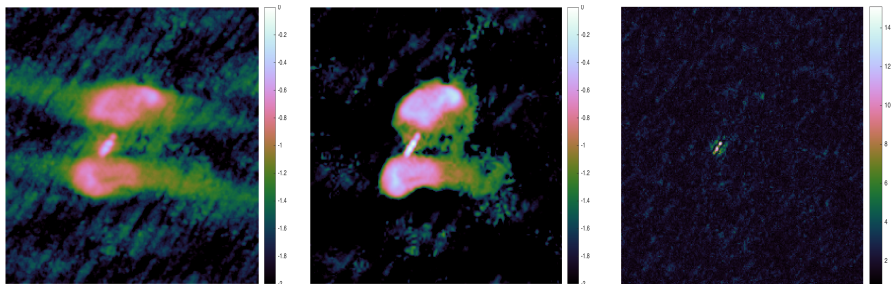
(a) Dirty image

(b) Mean recovered image

Figure: 3C288

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Figure: 3C288

Outline

- 1 Bayesian interpretations of interferometric imaging techniques
- 2 Proximal MCMC sampling
- 3 Sparse regularisation with Bayesian credible regions

Bayesian credible regions

Computation

- Combine **error estimation** with **fast sparse regularisation** (cf. compressive sensing) methods.
- Let C_α denote a **Bayesian credible region** with confidence level $(1 - \alpha)\%$:

$$P(\mathbf{x} \in C_\alpha | \mathbf{y}) = \int_{\mathbf{x} \in C_\alpha} p(\mathbf{x} | \mathbf{y}) d\mathbf{x} = 1 - \alpha$$

- Define C_α by posterior iso-contour:

$$C_\alpha := \{\mathbf{x} : g(\mathbf{x}) \leq \gamma_\alpha\}$$

- Analytic approximation of γ_α derived in Pereyra (2016b):

$$\tilde{\gamma}_\alpha = g_{\mathbf{y}}(\mathbf{x}_{\text{map}}) + N(\tau_\alpha + 1),$$

where $\tau_\alpha = \sqrt{16 \log(3/\alpha)/N}$ and $\alpha \in (4\exp(-N/3), 1)$.

- Compute \mathbf{x}_{map} by sparse regularisation and **estimate Bayesian credible regions**.

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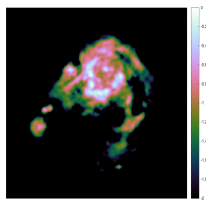
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Figure: HII region of M31

Bayesian credible regions

Preliminary results on simulations

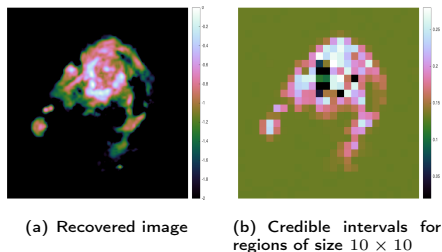
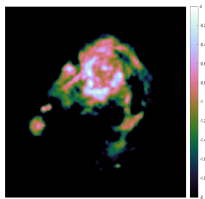


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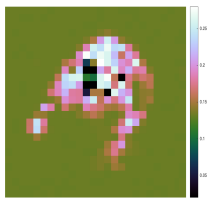
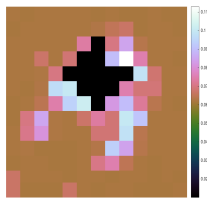
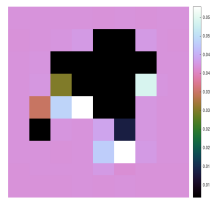
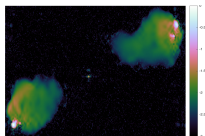
(b) Credible intervals for regions of size 10×10 (c) Credible intervals for regions of size 20×20 (d) Credible intervals for regions of size 30×30

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Bayesian credible regions

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Figure: Cygnus A

Bayesian credible regions

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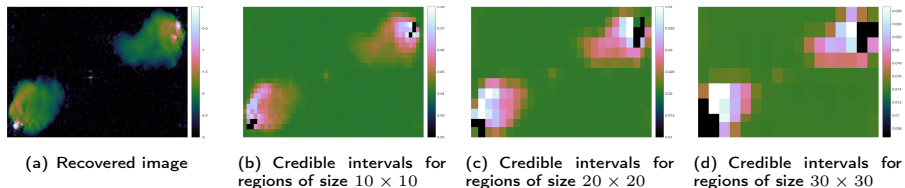
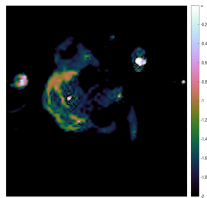


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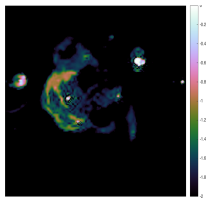


(a) Recovered image

Figure: Supernova remnant W28

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(a) Recovered image

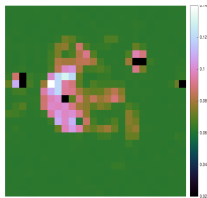
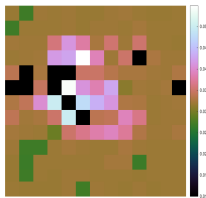
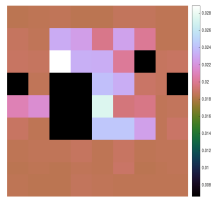
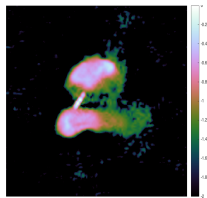
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Figure: Supernova remnant W28

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Preliminary results on simulations



(a) Recovered image

Figure: 3C288

Bayesian credible regions

Preliminary results on simulations

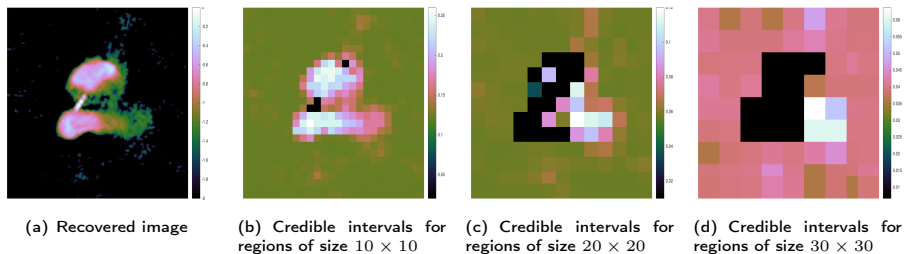


Figure: 3C288

Hypothesis testing

Method

- Is structure in an image **physical or an artefact**?
- Can we make **precise statistical statements**?
- Perform **hypothesis tests** using Bayesian credible regions (Pereyra 2016b).

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- Perform **hypothesis tests** using Bayesian credible regions (Pereyra 2016b).

Hypothesis testing of physical structure

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- 2 Inpaint background (noise) into region, yielding surrogate image \mathbf{x}'_{map} .
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 - If $\mathbf{x}'_{\text{map}} \notin C_\alpha$ then reject hypothesis that structure is an artefact with confidence $(1 - \alpha)\%$, i.e. structure most likely physical.
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Hypothesis testing

Method

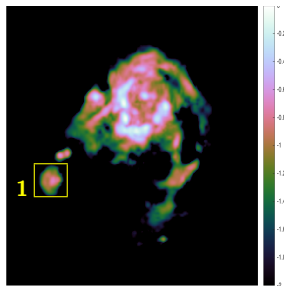
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Hypothesis testing

Preliminary results on simulations

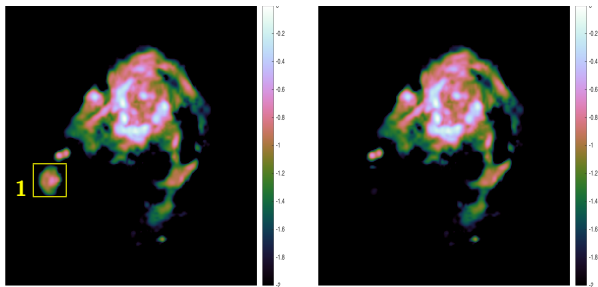


(a) Recovered image

Figure: HII region of M31

Hypothesis testing

Preliminary results on simulations



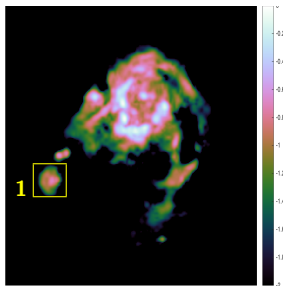
(a) Recovered image

(b) Surrogate with region removed

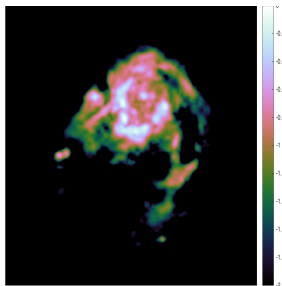
Figure: HII region of M31

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(a) Recovered image



(b) Surrogate with region removed

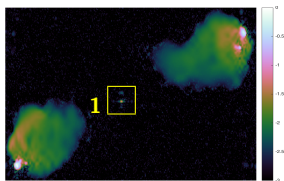
Reject null hypothesis

⇒ structure physical

Figure: HII region of M31

Hypothesis testing

Preliminary results on simulations



(a) Recovered image

Figure: Cygnus A

Hypothesis testing

Preliminary results on simulations

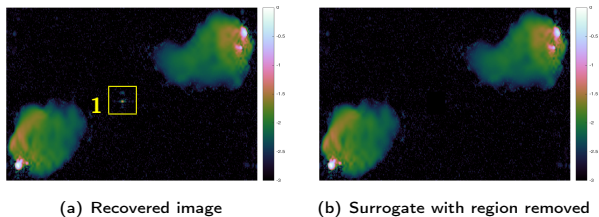
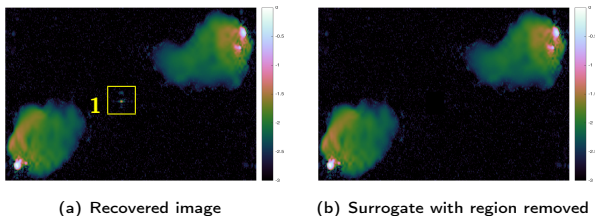


Figure: Cygnus A

Hypothesis testing

Preliminary results on simulations



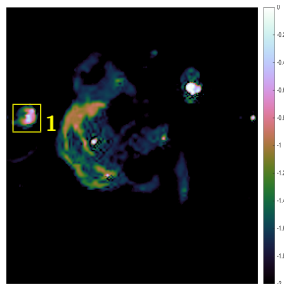
Cannot reject null hypothesis

⇒ cannot make strong statistical statement about origin of structure

Figure: Cygnus A

Hypothesis testing

Preliminary results on simulations



(a) Recovered image

Figure: Supernova remnant W28

Hypothesis testing

Preliminary results on simulations

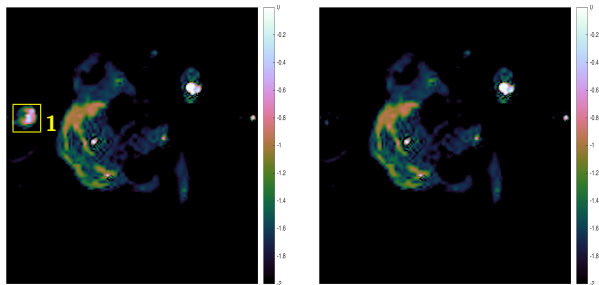
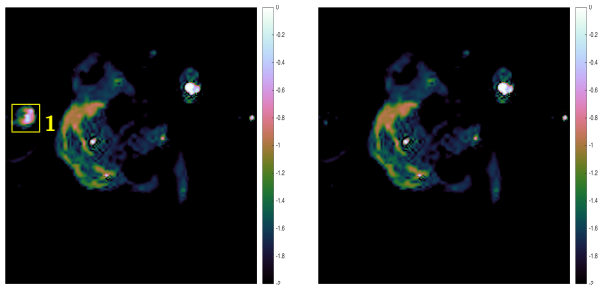


Figure: Supernova remnant W28

Hypothesis testing

Preliminary results on simulations



(a) Recovered image

(b) Surrogate with region removed

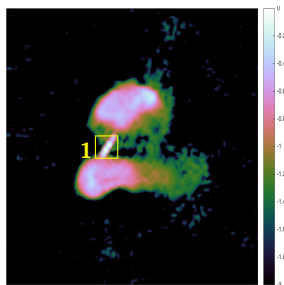
Reject null hypothesis

⇒ structure physical

Figure: Supernova remnant W28

Hypothesis testing

Preliminary results on simulations

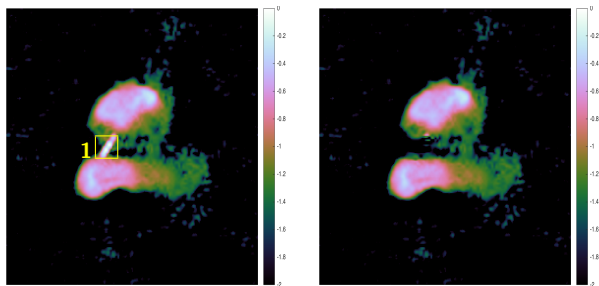


(a) Recovered image

Figure: 3C288

Hypothesis testing

Preliminary results on simulations



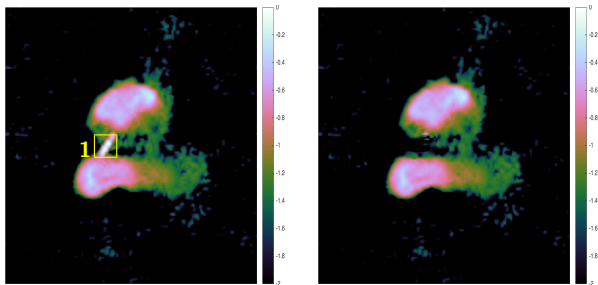
(a) Recovered image

(b) Surrogate with region removed

Figure: 3C288

Hypothesis testing

Preliminary results on simulations



(a) Recovered image

(b) Surrogate with region removed

Reject null hypothesis

⇒ structure physical

Figure: 3C288

Conclusions

- **Sparse priors** (cf. **compressive sensing**) shown to be highly effective (see talks by Luke Pratley, Alex Onose, Vijay Kartik).
- Also seek statistical interpretation to recover **error information**.
- **Proximal MCMC** sampling can support sparse priors in full statistical framework (P-MALA).
- Combine **error estimation** with **fast sparse regularisation**:
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